

Intrinsic Semiparametric and Nonparametric Models for Symmetric Positive Definite Matrices

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Outline

Medical Imaging

Regression Models for Symmetric Positive Definite Matrices

Nonparametric Models for Symmetric Positive Definite Matrices

Simulation Studies and Real Data Analysis

Future Work



Study function and development of brain functional and structural connectivity.









Euclidean-valued Data

intensity, fMRI, volume, grey matter density, SPHARM, invariant measure, signed-Euclidean distance, ...





Manifold-valued Data

Directional data, deformation tensors, diffusion tensors, principal directions, medial representation, projections, orientation, rigid motion,



Deformation Tensor

Diffusion Tensor



Principal Direction

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Medial Representation



White Matter Maturation





<figure>







Manifold-valued Data

 S_2^{\bullet}

Riemannian Response Space: Riemannian manifold is connected and geodesically complete

> ^C Euclidean Covariate Space

 $g(\bullet): \mathbb{R}^k \longrightarrow M$ Link Function

g(x)





Existing Literature in Statistics

Parametric/Semiparametric Statistical Inference in Euclidean Space: Rao (1945), Efron (1975), Amari (1985), Cook (1986), McCullagh (1987), Barndorff-Nielsen and Cox (1994), Wei (1988), ...

Statistics for Manifold-valued Data:

- <u>Directional Statistics:</u> Fisher (1953), Fisher (1993), Kent (1977), Watson (1983), Mardia and Jupp (1999), ...
- <u>Axial and Shape Spaces:</u> Kendall (1977, 1984), Dryden and Mardia (1998), Kendall, Barden, Carne, and Le (1999), ...
- **Diffusion Tensors:** Armin Schwartzman (2006, 2008), Fletcher and Joshi (2007), Dryden et al. (2009), Zhu et al. (2009), ...
- **<u>Riemmanian Manifold</u>**: Bhattacharya and Patrangenaru (2003a, b), ...
- **<u>Data Mining:</u>** Huckemann, S., Hotz, T., Munk, A. (2010), Huckemann et al. (2006), ..
- <u>Bayesian methods: Jermyn (2005),</u> Angers and Kim (2005), Bhattacharya and Dunson (2010), ...



Semiparametric and Nonparametric Regression for Manifold-valued Response from Cross-sectional, Longitudinal and Family-based Neuroimaging Studies





Symmetric Positive Definite Matrix (SPD)

- Diffusion Tensors in DTI are 3x3 SPDs. DTI is an imaging modality that allows measurement of fiber-tract trajectories in vivo in soft tissues.
- Covariance Matrices: Multivariate analysis, Longitudinal data, Spatial data, ...
- Network Data:







An appropriate statistical analysis of SPD matrices is important for understanding normal brain development, the neural bases of neuropsychiatric disorders, and the joint effects of environmental and genetic factors on brain structure and function.

Euclidean Space



A formal statistical framework for using a set of covariates in a Euclidean space to predict SPD matrices as responses:

- Extrinsic Methods: Ignore the fact that SPD matrices are in a nonlinear space and then directly apply classical multivariate regression. Schwartzman and Taylor (2008), ...
- Intrinsic Methods: Several parametric models for SPD matrices.





Dryden, I.L., Koloydenko, A. and Zhou, D. (2009).









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Scale Frobenius inner product

$$< Y_D, Z_D >= tr(Y_D D^{-1} Z_D D^{-1})$$

Geodesic

$$\gamma_D(t; Y_D) = G \exp(tG^{-1}Y_DG^{-T})G^T$$
$$D = GG^T$$

Riemannian exponential/logarithm maps

$$X = Exp_D(Y_D) = \gamma_D(1;Y_D) = G\exp(G^{-1}Y_DG^{-T})G^T$$
$$Y_D = Log_D(X) = G\log(G^{-1}XG^{-T})G^T$$





Riemannian logarithm map

$$Y_{\Sigma(x)} = Log_{\Sigma(x)}(S) = G(x)\log(G(x)^{-1}SG(x)^{-T})G(x)^{T}$$

$$\Sigma(x) = \Sigma(x,\theta) = G(x,\theta)G(x,\theta)^{T} = G(x)G(x)^{T}$$

Use Riemannian logarithm map to construct residuals

Rotate residuals to the same tangent plane (parallel transport)

Q1: Residual
$$e(x,\theta) = Log_{\Sigma(x,\theta)}(S) = \log(G(x,\theta)^{-1}SG(x,\theta)^{-T})$$



Q2: Link functions

$$\Sigma(x,\theta): R^k \times \Theta \to Sym(m)^+$$

Cholesky decomposition

$$\Sigma(x,\theta) = G(x,\theta)G(x,\theta)^{T}$$

$$G(x,\theta) = \begin{cases} g_{11}(x,\theta) & 0 & 0 \\ g_{21}(x,\theta) & g_{22}(x,\theta) & 0 \\ g_{31}(x,\theta) & g_{32}(x,\theta) & g_{33}(x,\theta) \end{cases}$$

 $g_{ii}(x,\!\theta) \!\geq\! 0$



Link functions

$$\Sigma(x,\theta): R^k \times \Theta \to Sym(m)^+$$

matrix logarithm link

$$\log(\Sigma(x,\theta)) = g(x,\theta)$$

geodesic link

 $\sum (x = 0, \theta) = D$ $Y_D(x, \theta) \bullet \sum (x, \theta)$ $\Sigma(x=0,\theta)=D,$ $\Sigma(x,\theta) = \gamma_D(t(x),Y_D(x,\theta)),$ $t(x)Y_D(x,\theta) = g(x,\theta)$

 $Sym(m)^+$

 $\gamma_D(t;Y_D(x,\theta))$



Q3: Conditional Moment Model

$$E[e(x,\theta) | x] = E[Log_{\Sigma(x,\theta)}(S) | x] = 0$$

Intrinsic least square estimator (ILSE)

$$\hat{\theta} = \arg\min G_n(\theta) = \arg\min \sum_{i=1}^n tr(Log_{\Sigma(x_i,\theta)}(S_i)Log_{\Sigma(x_i,\theta)}(S_i))$$

ILSE includes the intrinsic mean as a special case.

$$\sum_{i=1}^{n} tr(Log_{\Sigma(x_{i},\theta)}(S_{i})Log_{\Sigma(x_{i},\theta)}(S_{i})) = \sum_{i=1}^{n} d(S_{i},\Sigma(x_{i},\theta))^{2}$$



Annealing Optimization Algorithm

• Gradient algorithm for computing ILSE is relatively sensitive to the starting point.

$$\theta^{(r+1)} = \theta^{(r)} + \rho \{ -\nabla^2 G_n(\theta^{(r)}) \}^{-1} \nabla G_n(\theta^{(r)}) \qquad \theta \in \mathbb{R}^p$$

$$Hess\{G_n(\theta^{(r)})\}\delta_r = -\nabla G_n(\theta^{(r)}) \qquad \theta \in M$$
$$\theta^{(r+1)} = R_{\theta^{(r)}}(\delta_r)$$

• Gibbs sampler $\exp(-G_n(\theta)/\tau)$

Annealing Evolutionary Stochastic Approximation Monte Carlo Algorithm Liang (2010)



Estimation Theory

• Consistency
$$\hat{\theta} \xrightarrow{p} \theta_*$$

Asymptotic Normality

$$(E[\sum_{i=1}^{n} \{\partial_{\theta} tr(e(x_{i},\hat{\theta})^{2})\}^{\otimes 2}])^{-1/2} E\{-\nabla^{2} G_{n}(\hat{\theta})\}(\hat{\theta}-\theta_{*}) \xrightarrow{L} N(0,I_{p^{*}})$$



Testing Linear Hypothesis

$$H_0: A\theta = A_0$$
 v.s. $H_0: A\theta \neq A_0$

Wald/Score test statistics

Resampling method/false discovery rate to correct for multiple comparisons



Simulation Studies Cholesky decomposition $G(x,\theta) = \begin{pmatrix} x_i^T \beta_1 & 0 & 0 \\ x_i^T \beta_2 & x_i^T \beta_3 & 0 \\ x_i^T \beta_4 & x_i^T \beta_5 & x_i^T \beta_6 \end{pmatrix}$ $x_i = (1, z_i)^T$

Data model

$$S_i = G(x_i, \theta) \exp(E_i) G(x_i, \theta)^T$$
$$E_i \sim MN(0, \Omega)$$

Correlation

$$\Omega_1 = \begin{pmatrix} 0.6 & 0 & 0 \\ 0 & 0.6 & 0 \\ 0 & 0 & 0.6 \end{pmatrix} \qquad \qquad \Omega_2 = \begin{pmatrix} 0.6 & 0.3 & 0.3 \\ 0.3 & 0.6 & 0.3 \\ 0.3 & 0.3 & 0.6 \end{pmatrix}$$



Simulation Study I

Table 1. Bias (×10⁻²), RMS (×10⁻²), SE (×10⁻²), SD–SE (×10⁻²), and RS of all 12 parameters under Ω_1 and Ω_2 . BIAS denotes the bias of the mean of the ILSE estimates; RMS denotes the root-mean-square error; SE denotes the mean of the standard deviation estimates; SD–SE denotes the standard deviation of the standard deviation estimates; RS denotes the ratio of RMS over SD. Two different sample sizes {20, 80} and 500 simulated datasets were used for each case

			n = 20		n = 80					
	BIAS	RMS	SE	SD-SE	RS	BIAS	RMS	SE	SD-SE	RS
					Ω_1					
β_1	2.60	6.10	6.73	2.20	1.10	0.58	3.59	3.37	0.68	0.94
β_2	1.78	6.10	6.61	1.72	1.08	0.06	3.90	3.61	0.55	0.92
β3	1.88	7.06	6.96	1.32	0.98	0.69	3.91	3.56	0.35	0.91
β4	1.15	6.86	6.89	1.30	1.01	0.35	3.83	3.51	0.33	0.92
β5	3.83	15.34	17.24	4.22	1.12	1.08	8.35	8.58	1.20	1.02
β_6	2.83	15.07	16.97	3.76	1.12	0.54	8.45	8.46	0.95	1.01
β7	1.43	8.75	8.07	1.42	0.92	-0.37	4.19	4.10	0.39	0.98
β ₈	0.48	8.32	7.98	1.40	0.96	-0.44	4.12	4.06	0.38	0.98
β9	5.14	29.38	32.06	7.57	1.09	1.6	14.84	16.07	2.23	1.08
β10	3.97	28.88	31.59	7.14	1.09	1.05	14.84	15.85	1.77	1.07
β11	3.62	20.32	19.91	3.87	0.98	1.00	10.63	10.15	0.85	0.96
β_{12}	2.69	20.11	19.68	3.83	0.98	0.53	10.48	10.03	0.84	0.96
					Ω_2					
β_1	2.76	6.87	6.73	2.21	0.97	0.59	4.00	3.7	0.57	0.93
β_2	1.81	6.72	6.61	1.63	0.98	0.46	3.97	3.7	0.53	0.93
β ₃	1.96	7.74	7.23	1.27	0.93	0.23	3.72	3.54	0.36	0.95
β_4	1.24	7.43	7.15	1.26	0.96	0.03	3.67	3.5	0.36	0.95
β ₅	3.08	11.63	12.61	2.85	1.08	1.02	6.87	6.33	0.73	0.92
β ₆	1.87	11.78	12.41	2.31	1.05	0.86	6.9	6.3	0.7	0.91
β7	1.56	8.44	8.26	1.46	0.98	0.11	4.5	4.2	0.34	0.93
β_8	0.62	8.09	8.17	1.44	1.01	0.04	4.5	4.1	0.34	0.91
β9	2.75	18.90	19.75	4.14	1.04	1.06	10.74	9.93	1.09	0.93
β_{10}	1.33	18.84	19.46	3.68	1.03	0.86	10.79	9.82	0.93	0.91
β_{11}	3.87	16.68	15.49	2.54	0.93	-0.58	7.45	7.44	0.66	0.99
β12	2.88	16.46	15.31	2.51	0.94	-0.91	7.22	7.37	0.65	1.02



Simulation Study II

Table 2. Comparisons of the rejection rates for score test statistics under Ω_1 and Ω_2 . Three differing sample sizes {20, 40, 80} and 1,000 simulated datasets were used for each case and two significance levels, 5% and 1%, were considered. The null and alternative hypotheses were, respectively, given by $H_0: \boldsymbol{\beta}_{\cdot,2} = \mathbf{0}$ and $H_1: \boldsymbol{\beta}_{\cdot,2} \neq \mathbf{0}$. Two methods including the resampling method (RE) and χ^2 distribution [$\chi^2(6)$] were used to calculate the rejection rates

	n = 20				n = 40				n = 80			
	5%		1%		5%		1%		5%		1%	
β .,2	RE	χ ² (6)	RE	$\chi^{2}(6)$	RE	χ ² (6)	RE	χ ² (6)	RE	χ ² (6)	RE	χ ² (6)
						Ω_1						
$0 \times 1_{6}$	0.143	0.031	0.037	0	0.067	0.043	0.026	0.007	0.067	0.037	0.017	0.003
$0.2 \times 1_{6}$	0.513	0.177	0.253	0.011	0.957	0.883	0.796	0.461	1	1	0.991	0.971
$0.4 \times 1_{6}$	0.597	0.213	0.293	0.022	0.993	0.951	0.832	0.481	1	1	1	1
$0.6 \times 1_{6}$	0.773	0.442	0.520	0.042	1	1	1	0.983	1	1	1	1
						Ω2						
$0 \times 1_{6}$	0.126	0.023	0.037	0	0.063	0.037	0.017	0.003	0.061	0.033	0.013	0.003
$0.2 \times 1_{6}$	0.581	0.221	0.302	0.010	0.977	0.953	0.851	0.491	1	1	0.991	0.991
$0.4 \times 1_{6}$	0.602	0.227	0.321	0.032	0.991	0.981	0.871	0.562	1	1	1	1
$0.6 \times 1_{6}$	0.903	0.51	0.611	0.051	1	1	1	0.981	1	1	1	1
						-						



Simulation Study III

Table 3. Comparisons of the family-wise error rates and average powers for the test procedure under two different correlations $\rho = 0.0$ and 0.5. We considered two different sample sizes {40, 80} and 100 simulated datasets for each case at the 5% significance level. In all voxels, the null and alternative hypotheses were, respectively, given by $H_0: \beta_{.,2}(d) = \mathbf{0}_6$ and $H_1: \beta_{.,2}(d) \neq \mathbf{0}_6$. We considered four different $\beta_{.,2}(d)$ {0.0 × $\mathbf{1}_6, 0.3 \times \mathbf{1}_6, 0.6 \times \mathbf{1}_6, 0.9 \times \mathbf{1}_6$ } for all voxels within the region of interest, whereas we set $\beta_{.,2}(d) = \mathbf{0}_6$ for all voxels outside the region of interest. FWR denotes the family wise error rate and Apower denotes the average rejection rate for voxels inside the region of interest.

β .,2(d)		<i>n</i> =	= 40		n = 80				
	ρ	= 0.0	ρ	= 0.5	ρ	= 0.0	$\rho = 0.5$		
	FWR	Apower	FWR	Apower	FWR	Apower	FWR	Apower	
$0.0 \times 1_{6}$	0.12	0.00	0.06	0.00	0.08	0.00	0.07	0.00	
$0.3 \times 1_{6}$	0.18	0.10	0.12	0.10	0.06	0.56	0.06	0.57	
$0.6 \times 1_{6}$	0.14	0.67	0.06	0.68	0.02	1.00	0.03	1.00	
$0.9 \times 1_{6}^{\circ}$	0.12	0.83	0.10	0.85	0.08	1.00	0.06	1.00	



HIV Neuroimaging Data (PI: Colin Hall)

Objective: Assess diagnosis and age on the integrity of white matter in a cross-sectional study of human immunodeciency virus (HIV).

Participants: All 47 subjects with 29 HIV+ subjects (21 males and 8 females) and 18 healthy (9 males and 9 females) controls were studied. We limited the statistical analysis within the major white matter regions (FA>0.4).

Results: We observe statistically significant diagnosis effects in superior internal capsule area and age effects in inferior longitudinal fasciculus.



Group Effect







Data $(x_1, S_1), \dots, (x_n, S_n)$



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Styner, M. (2008).



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S(x)

arc-length

В

Α

E



$$T_{D}M$$

$$T_{D}M$$

$$T_{D}M$$

$$Y_{D}$$

$$Z_{D}$$

$$S = Exp_{D}(Y_{D})$$

$$\gamma_{D}(t; Y_{D})$$

 $M = Sym(m)^+$ Inner product $\langle \langle Y_D, Z_D \rangle \rangle$

Geodesic

Riemannian exponential/logarithm maps

Affine invariant metric

$$<< Y_D, Z_D >>_{D,R} = tr(Y_D D^{-1} Z_D D^{-1})$$

Log-Euclidean metric

$$<< Y_D, Z_D >>_{D,L} = tr(R_D(Y_D)R_D(Z_D))$$

$$R_D:T_DM\to T_{I_m}M$$

Dryden et al. (2009) Ying, Zhu, Lin and Marron (2010)



Questions:

- How to define <u>local polynomial kernel</u> regression to nonparametrically estimate an intrinsic mean of S given x?
- Whether <u>local linear</u> regression performs better than <u>local constant</u> regression?
- How much statistical inferences depend on <u>a specific inner product</u> defined on the tangent space?







Conditional Expectation

$$D(x) = E[S | X = x]$$

$$E[e(x) | X = x] = E[S - D(x) | X = x] = 0$$

Intrinsic Conditional Expectation

$$e_{D(x)} = \text{Log}_{D(x)}(S) \in T_{D(x)}Sym^{+}(m)$$
$$E[e_{D(x)} | X = x] = E[\log_{D(x)}(S) | X = x] = 0$$



Local Polynomial Kernel Regression

$$\begin{aligned} \text{Log}_{D(x_0)}(D(x)) &\in T_{D(x_0)} Sym^+(m) \\ \phi_{D(x_0)}(.) &: T_{D(x_0)} Sym^+(m) \to T_{I_m} Sym^+(m) \\ Y(x) &= \phi_{D(x_0)}(\text{Log}_{D(x_0)}(D(x))) \\ \text{Log}_{D(x_0)}(D(x))) &= \phi_{D(x_0)}^{-1}(Y(x)) \approx \phi_{D(x_0)}^{-1}(Y(x_0) + \sum_{k=1}^{K} Y^{(k)}(x_0)(x - x_0)^k) \\ D(x) &= \text{Exp}_{D(x_0)}(\phi_{D(x_0)}^{-1}(Y(x))) \approx \text{Exp}_{D(x_0)}(\phi_{D(x_0)}^{-1}(\sum_{k=1}^{K} Y^{(k)}(x_0)(x - x_0)^k)) \end{aligned}$$



Q1: Define Intrinsic LPK Estimator

$$\hat{\alpha}_{I}(x_{0}) = \operatorname{argmin}_{\alpha(x_{0})} \sum_{i=1}^{n} K_{h}(x_{i} - x_{0}) d(S_{i}, \operatorname{Exp}_{D(x_{0})}(\phi_{D(x_{0})}^{-1}(\sum_{k=1}^{K} Y^{(k)}(x_{0})(x - x_{0})^{k})))^{2}$$

K=0: Local constant estimator; *K=1*: Local linear estimator
Cross-validation

$$CV = \sum_{i=1}^{n} d(S_i, \hat{D}_I(x_i; h)^{(-i)})^2 \approx \sum_{i=1}^{n} d(S_i, \hat{D}_I(x_i; h))^2 + 2p_n(h)$$

Asymptotic average mean squared error (AMSE)

AMSE
$$(\log(\hat{D}_{IR}(x_0;h,k_0))) = E(tr[\{\log(\hat{D}_{IR}(x_0;h,k_0)) - \log(D(x_0))\}^2] | x)$$

Asymptotic average mean integrated squared error (AMISE)

$$AMISE(\log(\hat{D}_{IR}(h,k_0))) = \int AMISE(\log(\hat{D}_{IR}(x;h,k_0)))\omega(x)dx$$



Log-Euclidean metric

$$\langle \langle U_{D(x)}, V_{D(x)} \rangle \rangle_{D(x),L} = \operatorname{tr}(R_{D(x)}(U_{D(x)})R_{D(x)}(V_{D(x)}))$$

Let

$$Y(x) = \phi_{D(x_0),L}(\text{Log}_{D(x_0)}(D(x))) = \log(D(x)) - \log(D(x_0))$$

$$D(x) = \exp(\log(D(x_0)) + Y(x))$$

Intrinsic Mean

$$E[\log(S) - \log(D(x)) | X = x] = O_m$$

Geodesic Distance

$$d(D(x),S)^2 = tr\{[log(D(x)) - log(S)]^2\}$$

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Q2: Intrinsic local linear is better than intrinsic local constant AMSE $(\log(\hat{D}_{u}(x_{0};h,0))) = h^{4}u_{2}^{2}tr([vecs\{0.5 \times \log(D(x_{0}))^{(2)} + f_{x}^{(1)}(x_{0})f_{x}(x_{0})^{-1}\log(D(x_{0}))^{(1)}\}]^{\otimes 2})$ $+ v_0 (nhf_x(x_0))^{-1} tr(\Sigma_{sD}(x_0)) + o(h^4 + (nh)^{-1})$ $h_{opt,L}(x_0;0) = \left[\frac{n^{-1}v_0 f_X^{-1}(x_0) \operatorname{tr}(\Sigma_{\varepsilon D}(x_0))}{4u_2^2 \operatorname{tr}([\operatorname{vecs}\{0.5\log(D(x_0))^{(2)} + f_X^{(1)}(x_0) f_X(x_0)^{-1}\log(D(x_0))^{(1)}\}]^{\otimes 2})}\right]$ **Optimal bandwidth** AMSE $(\log(\hat{D}_{u}(x_{0};h,1))) = h^{4}u_{2}^{2}tr([vecs\{0.5 \times \log(D(x_{0}))^{(2)}\}]^{\otimes 2}) + v_{0}(nhf_{x}(x_{0}))^{-1}tr(\Sigma_{eD}(x_{0}))$ $+ o(h^4 + (nh)^{-1})$ $h_{opt,L}(x_0;1) = \left| \frac{n^{-1} v_0 f_X^{-1}(x_0) \operatorname{tr}(\Sigma_{\varepsilon D}(x_0))}{u_2^2 \operatorname{tr}([\operatorname{vecs}\{\log(D(x_0))^{(2)}\}]^{\otimes 2})} \right|$ **Optimal bandwidth Ratio of AMSEs** $\frac{\text{AMSE}(\log(\hat{D}_{IL}(x_0;h,0)))}{\text{AMSE}(\log(\hat{D}_{IL}(x_0;h,1)))} = \frac{\text{tr}([\text{vecs}\{0.5\log(D(x_0))^{(2)} + f_X^{(1)}(x_0)f_X(x_0)^{-1}\log(D(x_0))^{(1)}\}]^{\otimes 2})}{\text{tr}([\text{vecs}\{0.5\log(D(x_0))^{(2)}\}]^{\otimes 2})}$



Affine invariant metric

$$<< U_{D(x)}, V_{D(x)} >>_{D(x),R} = tr(U_{D(x)}D(x)^{-1}V_{D(x)}D(x)^{-1})$$

Let
$$D(x) = G(x)G(x)^{T}$$

 $Y(x) = \phi_{D(x_{0}),R}(\text{Log}_{D(x_{0})}(D(x))) = \log(G(x_{0})^{-1}D(x)G(x_{0})^{-1})$
 $D(x) = G(x_{0})\exp(Y(x))G(x_{0})^{T}$

Intrinsic Mean

$$E[\log(G(x)^{-1}SG(x)^{-1}) | X = x] = O_m$$

Geodesic distance

$$d(D(x),S)^{2} = \operatorname{tr}\{\log^{2}(G(x)^{-1}SG(x)^{-1})\}$$

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Q2: Intrinsic local linear is better than intrinsic local constant

$$\begin{aligned} AMSE(\log(\hat{D}_{IR}(x_{0};h,0))) &= h^{4}u_{2}^{2} \text{tr}([\mathbf{G}_{D}(x_{0})^{T} \operatorname{vecs}\{\mathbf{G}^{(1)}(x_{0})f_{X}^{(1)}(x_{0})f_{X}(x_{0})^{-1} + 0.5\mathbf{G}^{(2)}(x_{0})\}]^{\otimes 2}) \\ &+ (nh)^{-1} \text{tr}(\mathbf{G}_{D}(x_{0})^{\otimes 2}\Omega_{0}(x_{0})) + o(h^{4} + (nh)^{-1}) \end{aligned}$$

$$\begin{aligned} \mathbf{Optimal\ bandwidth} \ h_{opt,R}(x_{0};0) &= \left[\frac{n^{-1}\text{tr}(G_{D}(x_{0})^{\otimes 2}\Omega_{0}(x_{0}))}{4u_{2}^{2}\text{tr}([\mathbf{G}_{D}(x_{0})^{T} \operatorname{vecs}\{G^{(1)}(x_{0})f_{X}^{(1)}(x_{0})f_{X}(x_{0})^{-1} + 0.5G^{(2)}(x_{0})\}]^{\otimes 2})} \right]^{1/5} \\ AMSE(\log(\hat{D}_{IR}(x_{0};h,1))) &= 0.25h^{4}u_{2}^{2}\text{tr}([\mathbf{G}_{D}(x_{0})^{T}\Psi_{1}(x_{0})^{-1}\Psi_{2}(x_{0})^{T} \operatorname{vecs}\{\mathbf{Y}^{(2)}(x_{0})\}]^{\otimes 2}) \\ &+ (nh)^{-1}\text{tr}(\mathbf{G}_{D}(x_{0})^{\otimes 2}\Omega_{0}(x_{0})) + o(h^{4} + (nh)^{-1}) \end{aligned}$$

$$\begin{aligned} \mathbf{Optimal\ bandwidth} \ h_{opt,R}(x_{0};1) &= \left[\frac{n^{-1}\text{tr}(G_{D}(x_{0})^{\otimes 2}\Omega_{0}(x_{0}))}{4u_{2}^{2}\text{tr}([\mathbf{G}_{D}(x_{0})^{T}\Psi_{1}(x_{0})^{-1}\Psi_{2}(x_{0})^{T} \operatorname{vecs}\{Y^{(2)}(x_{0})\}]^{\otimes 2})} \right]^{1/5} \end{aligned}$$

Ratio of AMSEs

$$\frac{\text{AMSE}(\log(\hat{D}_{IR}(x_0;h,0)))}{\text{AMSE}(\log(\hat{D}_{IR}(x_0;h,1)))} = \frac{\text{tr}([G_D(x_0)^T \operatorname{vecs}\{G^{(1)}(x_0)f_X^{(1)}(x_0)f_X^{(1)}(x_0)f_X^{(1)}(x_0)^{-1} + 0.5G^{(2)}(x_0)\}]^{\otimes 2})}{\text{tr}([G_D(x_0)^T \Psi_1(x_0)^{-1} \Psi_2(x_0)^T \operatorname{vecs}\{Y^{(2)}(x_0)\}]^{\otimes 2})}$$



Q3: Affine Invariant Metric versus Log-Euclidean Metric

$$rMSE(R,L;0) = \frac{AMSE(log(\hat{D}_{IR}(x_0;h,0)))}{AMSE(log(\hat{D}_{IL}(x_0;h,0)))} = \left[\frac{tr\{G_D(x_0)^{\otimes 2}\Psi_1(x_0)^{-1}\Psi_{11}(x_0)\Psi_1(x_0)^{-1}\}}{tr\{\Sigma_{eD}(x_0)\}}\right]^{4/5} \\ \times \left[\frac{tr([G_D(x_0)^T vecs\{f_X^{(1)}(x_0)f_X(x_0)^{-1}G^{(1)}(x_0)+0.5G^{(2)}(x_0)\}]^{\otimes 2})}{tr([vecs\{0.5log(D(x_0))^{(2)}+f_X^{(1)}(x_0)f_X(x_0)^{-1}log(D(x_0))^{(1)}\}]^{\otimes 2})}\right]^{1/5}$$

Let m=1,
$$D(x_0) = G(x_0)^2$$
 with $G(x_0) > 0$
rMSE(*R*,*L*;0) = $\left[\frac{\operatorname{tr}([f_X^{(1)}(x_0)f_X(x_0)^{-1}G^{(1)}(x_0) + 0.5G^{(2)}(x_0)]^{\otimes 2})}{\operatorname{tr}([f_X^{(1)}(x_0)f_X(x_0)^{-1}G^{(1)}(x_0) + 0.5G^{(2)}(x_0) - 0.5G^{(1)}(x_0)^2G(x_0)^{-1}]^{\otimes 2})}\right]^{1/5}$

 $f_X^{(1)}(x_0)f_X(x_0)^{-1}G^{(1)}(x_0) + 0.5G^{(2)}(x_0) > 0.25G^{(1)}(x_0)^2G(x_0)^{-1} \Leftrightarrow \mathrm{rMSE}(R,L;0) > 1$

It depends on both design density and $D(x_0) = G(x_0)^2$.



Q3: Affine Invariant Metric versus Log-Euclidean Metric

$$rMSE(R,L;1) = \left[\frac{tr\{G_D(x_0)^{\otimes 2}\Omega_0(x_0)\}}{tr\{\Sigma_{eD}(x_0)\}}\right]^{4/5}$$

$$\times \left[\frac{tr([G_D(x_0)^T\Psi_1(x_0)^{-1}\Psi_2(x_0)^T \operatorname{vecs}(Y^{(2)}(x_0))\}^{\otimes 2})}{tr(\operatorname{vecs}\{\log(D(x_0))^{(2)}\}^{\otimes 2})}\right]^{1/5}$$
Let m=1, $D(x_0) = G(x_0)^2$ with $G(x_0) > 0$
 $rMSE(R,L;1) = \frac{AMSE(\log(\hat{D}_{IR}(x_0;h,1)))}{AMSE(\log(\hat{D}_{IL}(x_0;h,1)))} = 1$



Simulation Studies

Data model $S_{i} = C(x_{i}) \exp(E_{i})C(x_{i}), \quad E_{i} \sim MN(0,\Omega)$ $D(x) = C(x)^{2} \qquad x_{i} \sim N(0,0.25)$ $C(x) = \begin{pmatrix} -0.1x & 0.2x & \sin(x) \\ 0.2x & 0.6x & -0.4x \\ \sin(x) & -0.4x & 0.5x \end{pmatrix}$

Covariance

$$\Sigma_{1} = \begin{pmatrix} 0.3 & 0.049 & 0.052 \\ 0.049 & 0.2 & 0.0424 \\ 0.052 & 0.0424 & 0.1 \end{pmatrix} \qquad \Sigma_{2} = 2\Sigma_{1}, \qquad \Sigma_{3} = 4\Sigma_{1}, \qquad \Sigma_{4} = 8\Sigma_{1}$$

<u>Data</u> $\{(x_i, S_i) : i = 1, \dots, n\}$ for n = 50 or 100



$$3 \times 3 \quad SPD \qquad D(x) > 0 \qquad \lambda_1 \ge \lambda_2 \ge \lambda_3 > 0$$
$$FA = \sqrt{\frac{3\{(\lambda_1 - \overline{\lambda})^2 + (\lambda_2 - \overline{\lambda})^2 + (\lambda_3 - \overline{\lambda})^2\}}{2(\lambda_1^2 + \lambda_2^2 + \lambda_3^2)}}$$

FA: a scalar quantity derived from diffusion tensor (SPD matrix).

- Low FA values: isotropic diffusion.
- High FA values: highly directional diffusion.







Fig. 1. Ellipsoidal representations of the true (the first row) and simulated SPD matrix data along the design points under the four different noise distributions (the second to the fifth rows: Σ_1 - Σ_4) colored with FA values.









Simulation 1.

Compare the performance of the local linear with the local constant

• Assess the performance using the Average Geodesic Distance (AGD) for each replication j=1, ..., N with N as the number of replications, denoted by N_n

AGD =
$$(nN)^{-1} \sum_{j=1}^{N} \sum_{i=1}^{n} d(\hat{D}_{j}(x_{i}), D(x_{i}))$$

where $\hat{D}_j(x_i)$ and $D(x_i)$ diffusion tensors at X_i

and $D(x_i)$ are, respectively, the estimated and true





Fig. 3. Boxplots of the AGD using the intrinsic local constant and linear estimators under the log-Euclidean (the first row) and Riemannian (the second row) metrics based on 100 replications under the three covariance matrices (a)-(b) Σ_1 , (c)-(d) Σ_2 , and (e)-(f) Σ_3 . C50 and C100 represent the intrinsic local constant estimators at sample sizes 50 and 100, respectively. L50 and L100 represent the intrinsic local linear estimators at sample sizes 50 and 100, respectively.





Local linear (dashed) Fig. 4. The LAGD curves at each sample point using the intrinsic local constant (solid line) and linear (dash-dotted line) estimators under the three covariance matrices (a)-(d) Σ_1 , (e)-(h) Σ_2 , (i)-(l) Σ_3 for sample sizes 50 (the top two rows) and 100 (the bottom two rows). The first and third rows correspond to the log-Euclidean metric while the second and fourth rows correspond to the Riemannian metric.



Simulation 2. High noisy level Compare the performance of the local linear under two metrics



Fig. 5. (a) Boxplots of the AGD's using the linear regressions based on 100 replications under the covariance matrix Σ_4 , under the Log-Euclidean and Riemannian metrics, respectively. (b) and (c) LAGD curves at each sample point using the local linear regressions under the affine invariant (dash-dotted line) and Log-Euclidean (solid line) metrics under the the covariance matrix Σ_4 at sample size 50 (b) and 100 (c), respectively. LL50 (LR50) and LL100 (LR100), respectively, represent the local linear regressions under Log-Euclidean (Riemannian) metrics at sample sizes 50 and 100.



Simulation 3.

- Value of developing the LPK smoothing method
- Two different methods for smoothing FA values

M1. Calculate FA values from `noisy' SPDs and then use the local linear method to smooth the FA values

M2. Use the local linear method to smooth SPDs and then calculate FA values from the smoothed SPDs

• Calculate the Mean Absolute Deviation Error (MADE):

MADE =
$$(nN)^{-1} \sum_{j=1}^{N} \sum_{i=1}^{n} |FA_{j}(x_{i}) - FA_{j}(x_{i})|$$





Fig. 6. Boxplot of the MADE's using the two smoothing methods based on 100 replications under the covariance matrices (a) Σ_1 , (b) Σ_2 , and (c) Σ_3 at sample size 50. Smoothed FA curves for the realizations with median MADE under the covariance matrices: (d) Σ_1 , (e) Σ_2 , and (f) Σ_3 . The true FA curve (the solid line), the estimated FA curve using the first method (the dash-dotted line) and the estimated FA curve using the second method (the dashed line). This shows that the more intrinsic approach is much better.





Fig. 7. (a)The splenium of the corpus callosum in the analysis of HIV DTI data. (b)The ellipsoidal representation of full tensors on the fiber tract from a selected subject.



Fig. 8. (a) Ellipsoidal representations of the diffusion tensor data and estimated tensors using the intrinsic local linear regression under the (b)log-Euclidean and (c) Riemannian metrics along the fiber tract f1 colored with FA values. The estimated tensors in the middle right part (highlighted in the red line) are more anisotropic using the method under the Log-Euclidean metric.





Fig. 9. (a) FA's , (b) MD's and (c) PE's derived from the raw tensor data (dot line) and estimated tensors using the intrinsic local linear regression under the Riemannian (dash-dot line) and log-Euclidean (dash line) metrics as the function of arc-length along the tract f1. Estimated FA function along the fiber tract f1 by using the standard local linear regression for scalars (solid line).





Fig. 10. Ellipsoidal representations of estimated mean tensors along the fiber tract f1 for the control and HIV groups using the intrinsic local linear regression under the log-Euclidean ((a) and (b)) and Riemannian ((c) and (d)) metrics colored with FA values.





Fig. 11. (a) FA differences and (b) geodesic distances between pairs of mean diffusion tensors of HIV and control groups along the fiber f1 under the Log-Euclidean (the solid line) and Riemannian (the dashed line) metrics.



Manifold Data

• How to characterize the `variation' in the manifold data and use such information for statistical analysis?

• For a specific type of manifold data, is it possible to define an optimal geometrical structure that can capture the most important target information from the data at hand?

How to efficiently analyze discrete time and continuous time manifold data?



Future Work

- Minimax efficiency for the local linear method for SPD
- Statistical models for Lie group
- Statistical models of correlated Manifold-valued data
- Statistical models for deformation field
- Statistical models for spherical needlets



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