

Multilevel Modeling Day 2

Intermediate and Advanced Issues: Multilevel Models as Mixed Models

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What are mixed models

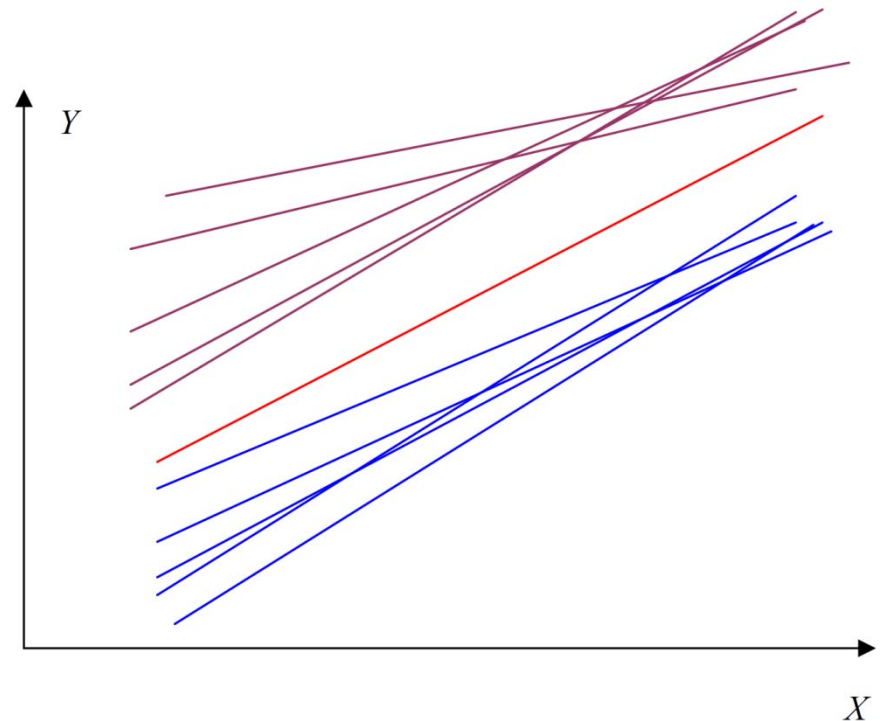
- The simplest multilevel models are in fact mixed models: have fixed and random parameters or effects

(3.3)
$$Y_{ij} = \gamma_{00} + u_{0j} + r_{ij} .$$

- γ_{00} is a fixed parameter and τ_{00} and σ^2 are random parameters
- The theory of mixed models is directly applicable
- Mixed effect models: regression analysis models with two types of effects:
 - Fixed effects (intercepts, slopes): describe the population studied as a whole. These effects are just like intercepts and slopes in conventional regression models.
 - Random effects: “bumps” up and down on the population intercepts and slopes, which are used to describe subpopulations. These effects can vary across subpopulations

Illustration of the Fixed and Random Effects

- X-socioeconomic status (SES); Y-achievement
- 10 schools
- Red line: population relationship line
 - describes an overall pattern of relationship
 - A “stable” relationship, i.e., its intercept and slope are fixed.
- Within each school, the intercepts and slopes vary
 - Typically schools are randomly drawn
 - Random effects



Mixed Model

- ❑ Mixed model are used for multilevel modeling, growth curve analysis, panel analysis or cross-sectional time series analysis
- ❑ Multilevel modeling: GPA related to study hours, but this relationship may differ by major
 - Fixed effects: overall relationship that GPA increases with study hours
 - Random effects: for a given major, the way this relationship differs from the overall one
- ❑ Longitudinal modeling: fluid intelligence declines with age in different nursing homes
 - Fixed effects: general trend
 - Random effects: for a given nursing home, it may have its own intercept and slope parameters. The amounts by which they differ from those in the overall population are represented by the random effects.
- ❑ More complex models: further nesting
- ❑ STATA comments:
 - xtmixed: fit mixed regression models
 - xtmelogit: fit mixed models with binary outcomes
 - xtmepoisson: fit mixed models to count data

Data for Demonstration

- ❑ Percentage voted for Bush (*bush*)
- ❑ Logarithm of number of people living in a square mile (*logdens*)
- ❑ Population minority (*minority*)
- ❑ Proportion adults aged 25 or higher with at least 4 years of college education (*colled*)
- ❑ Census division—out of 9 geographic regions (*cendiv*)

```
. use c:\stata\data\presidential_elections_2004.dta, clear
```

```
. d
```

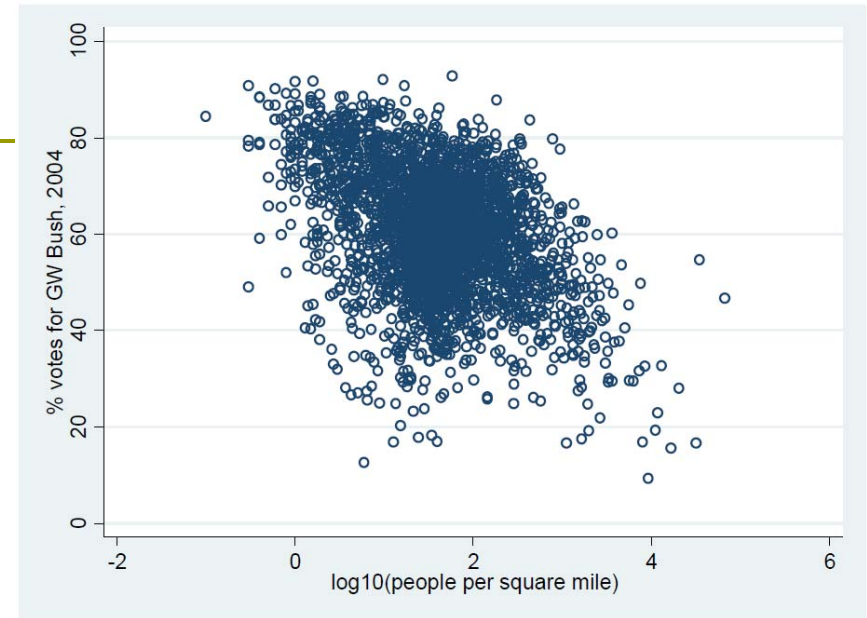
```
Contains data from c:\stata\data\presidential_elections_2004.dta
  obs:           3,054                US counties -- 2004 election (Robinson,
2005)
  vars:           11
  size:          232,104 (97.8% of memory free)      6 May 2011 15:21
```

variable name	storage type	display format	value label	variable label
fips	long	%9.0g		FIPS code
state	str20	%20s		State name
state2	str2	%9s		State 2-letter abbreviation
region	byte	%9.0g	region	Region (4)
cendiv	byte	%15.0g	division	Census division (9)
county	str24	%24s		County name
votes	float	%9.0g		Total # of votes cast, 2004
bush	float	%9.0g		% votes for GW Bush, 2004
logdens	float	%9.0g		log10(people per square mile)
minority	float	%9.0g		% population minority
colled	float	%9.0g		% adults >25 w/4+ years college

```
Sorted by:  fips
```

NOT Mixed Model

- ❑ Misleading: pools together politically, geographically, and economically diverse counties.
- ❑ There may be important differences across them with respect to the relationship between percentage voted for Bush and other characteristics (urban-rural differential)
- ❑ The traditional modeling approach does not capture the geographic differences in voting.
- ❑ The conventional approach assumes the same intercept and slope for all 3054 counties



Source	SS	df	MS	Number of obs = 3041		
Model	122345.617	3	40781.8725	F(3, 3037) = 345.39		
Residual	358593.826	3037	118.075017	Prob > F = 0.0000		
Total	480939.443	3040	158.203764	R-squared = 0.2544		
				Adj R-squared = 0.2537		
				Root MSE = 10.866		

bush	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
logdens	-5.457462	.3031091	-18.00	0.000	-6.051781	-4.863142
minority	-.251151	.0125261	-20.05	0.000	-.2757115	-.2265905
colled	-.1811345	.0334151	-5.42	0.000	-.246653	-.115616
_cons	75.78636	.5739508	132.04	0.000	74.66099	76.91173

Conventional Regression -> Random Intercept

- The fitted model:

$$(2.1) \quad y_i = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \varepsilon_i, \quad (i = 1, \dots, n).$$

- where X_1 =logdens, X_2 =minority, X_3 =colled
- The subindex signals individual county
- Right-hand side says nothing about possible geographic differences
- None of the β s is county specific but assumed to be the same for all counties.

- Consider different geographic regions (9 census divisions: e.g., New England, Middle Atlantic, Mountain, Pacific).

- Allow each region to have its own intercept: random intercept model
- Each region is permitted to have its own intercept, but there's nothing more specific to it.
- u_{0j} is the unique region effect in the model
- The mean percent voting for Bush (mean of y) across counties within region is not the same in all regions at the average predictor values.

$$(2.2) \quad y_{ij} = \beta_0 + u_{0j} + \beta_1 X_{1,ij} + \beta_2 X_{2,ij} + \beta_3 X_{3,ij} + \varepsilon_{ij}$$
$$= \beta_{0j} + \beta_1 X_{1,ij} + \beta_2 X_{2,ij} + \beta_3 X_{3,ij} + \varepsilon_{ij} .$$

Random Intercept Model

```
. xtmixed bush logdens minority colled || cendiv:
```

Performing EM optimization:

Performing gradient-based optimization:

Iteration 0: log likelihood = -11339.79
Iteration 1: log likelihood = -11339.79

Computing standard errors:

Mixed-effects ML regression	Number of obs	=	3041
Group variable: cendiv	Number of groups	=	9
	Obs per group: min	=	67
	avg	=	337.9
	max	=	616
Log likelihood = -11339.79	Wald chi2(3)	=	1161.96
	Prob > chi2	=	0.0000

bush	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
logdens	-4.52417	.3621775	-12.49	0.000	-5.234025 -3.814316
minority	-.3645394	.0129918	-28.06	0.000	-.3900029 -.3390758
colled	-.0583942	.0357717	-1.63	0.103	-.1285053 .011717
_cons	72.09305	2.294038	31.43	0.000	67.59682 76.58929

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]
cendiv: Identity			
sd(_cons)	6.617132	1.600465	4.119005 10.63034
sd(Residual)	10.00339	.1284657	9.754742 10.25837

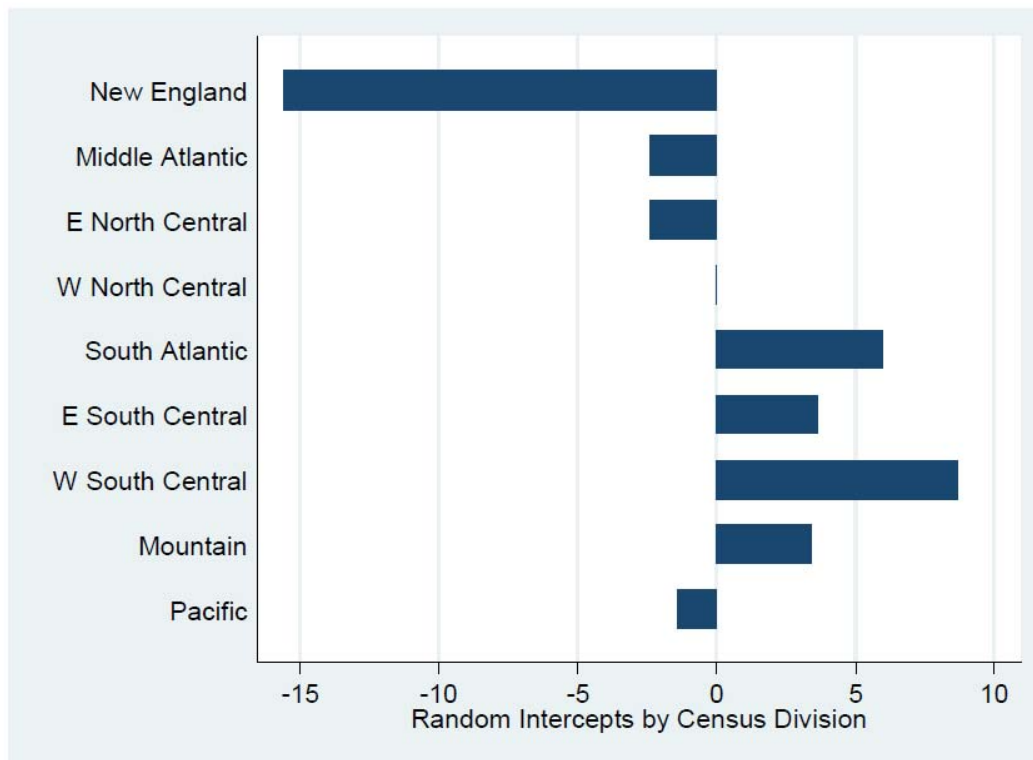
LR test vs. linear regression: chibar2(01) = 455.99 Prob >= chibar2 = 0.0000

- LR test: the RIM is a better model than the traditional regression
- The standard deviation of the random intercepts is found to be significant
- (2.4) gives the estimated model. Stars indicate that it is an estimated model

$$(2.4) \text{ bush}_{ij} = 72.13 - 4.70 \log \text{dens}_{ij} - .37 \text{ minority}_{ij} - .04 \text{ colled}_{ij} + *u_{0j} + *\epsilon_{ij}$$

Random Intercept Model

```
. predict randint0, reffects  
. graph hbar (mean) randint0, over(cendiv)
```



- Evaluate the regional intercepts
- Produce best linear unbiased predictions (BLUPs) of the random effects
- randint0 is a new variable with values for 3054 counties from 9 regions, it has the same value within each region. The value is actually the mean of all its values for a given region
- Interpretation: at any given level of the 3 predictors (logdens, minority and colled), percentage of voters for Bush is on average 16 points lower in the New England (NE) region than in the West/North/Central region, and 22 percent lower in the NE region than in the South Atlantic region.

Random Intercept Model

- List all the BLUP values for the random effects: the same values were assigned to each county in the same census division.
- BLUP is used in linear mixed models for the estimation of random effects. It is similar to the best linear unbiased estimates (BLUEs) of fixed effects.
 - Robinson, G.K. (1991). "That BLUP is a Good Thing: The Estimation of Random Effects". *Statistical Science* 6(1): 15–32
 - [Henderson, C.R.](#) (1975). "Best linear unbiased estimation and prediction under a selection model". *Biometrics* 31 (2): 423–447.

First consider an easy situation. Assumed model is a one-way random effects model with known intercept and variance components and normally distributed random effects:

$$Y_{it} = \mu + b_i + \varepsilon_{it}, t = 1, \dots, n_i; i = 1, \dots, q$$

$$b_i \sim i.i.d. N(0, \sigma_b^2)$$

$$\varepsilon_{it} \sim i.i.d. N(0, \sigma_\varepsilon^2)$$

$$\varepsilon_{it} \perp b_i, \mu, \sigma_\varepsilon^2, \text{ and } \sigma_b^2 \text{ known}$$

In which case the Best Linear Unbiased Predictor is given by

$$\tilde{b}_i = \frac{\sigma_b^2}{\sigma_b^2 + \sigma_\varepsilon^2 / n_i} (\bar{Y}_{i\cdot} - \mu)$$

. list cendiv randint0 in 1/100

	cendiv	randint0
1.	New England	-15.57527
2.	New England	-15.57527
3.	New England	-15.57527
4.	New England	-15.57527
5.	New England	-15.57527
6.	New England	-15.57527
7.	New England	-15.57527
8.	New England	-15.57527
9.	New England	-15.57527
10.	New England	-15.57527
11.	New England	-15.57527
12.	New England	-15.57527
13.	New England	-15.57527
14.	New England	-15.57527
15.	New England	-15.57527
16.	New England	-15.57527
17.	New England	-15.57527
18.	New England	-15.57527
19.	New England	-15.57527
20.	New England	-15.57527
21.	New England	-15.57527
22.	New England	-15.57527
23.	New England	-15.57527
24.	New England	-15.57527
25.	New England	-15.57527
67.	New England	-15.57527
68.	Middle Atlantic	-2.395582
69.	Middle Atlantic	-2.395582
70.	Middle Atlantic	-2.395582
71.	Middle Atlantic	-2.395582
72.	Middle Atlantic	-2.395582
73.	Middle Atlantic	-2.395582
74.	Middle Atlantic	-2.395582
75.	Middle Atlantic	-2.395582
76.	Middle Atlantic	-2.395582
77.	Middle Atlantic	-2.395582
78.	Middle Atlantic	-2.395582
79.	Middle Atlantic	-2.395582

Random Intercept Model

```
. xtmixed bush || cendiv:
```

```
Performing EM optimization:
```

```
Performing gradient-based optimization:
```

```
Iteration 0: log restricted-likelihood = -11895.132
Iteration 1: log restricted-likelihood = -11895.132
```

```
Computing standard errors:
```

```
Mixed-effects REML regression      Number of obs      =    3054
Group variable: cendiv              Number of groups    =      9

                                   Obs per group: min =     67
                                   avg =    339.3
                                   max =    618
```

```
Log restricted-likelihood = -11895.132      Wald chi2(0)      =      .
                                           Prob > chi2      =      .
```

bush	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
_cons	58.23352	2.383564	24.43	0.000	53.56182 62.90522

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]
cendiv: Identity			
sd(_cons)	7.103931	1.825418	4.293139 11.755
sd(Residual)	11.82181	.1514928	11.52859 12.12249

```
LR test vs. linear regression: chibar2(01) = 368.34 Prob >= chibar2 = 0.0000
```

```
. predict randint1, reffects
```

```
. table cendiv, contents (mean randint0 mean randint1 mean bush)
```

Census division (9)	mean(randint0)	mean(randint1)	mean(bush)
New England	-15.57527	-15.22133	42.38305
Middle Atlantic	-2.395582	-4.124419	54.03296
E North Central	-2.391467	-.4009925	57.82999
W North Central	.0169212	5.126793	63.38329
South Atlantic	5.96793	.8395588	59.07743
E South Central	3.617538	1.77551	60.02254
W South Central	8.68867	7.416742	65.69396
Mountain	3.456139	7.486413	65.79398
Pacific	-1.384879	-2.898272	55.2749

Can we use the formula from last slide to calculate the BLUP?

Consider New England:

$\sigma_b^2 = 7.103931^2$

$\sigma_e^2 = 11.82181^2$

$n_i = 67$

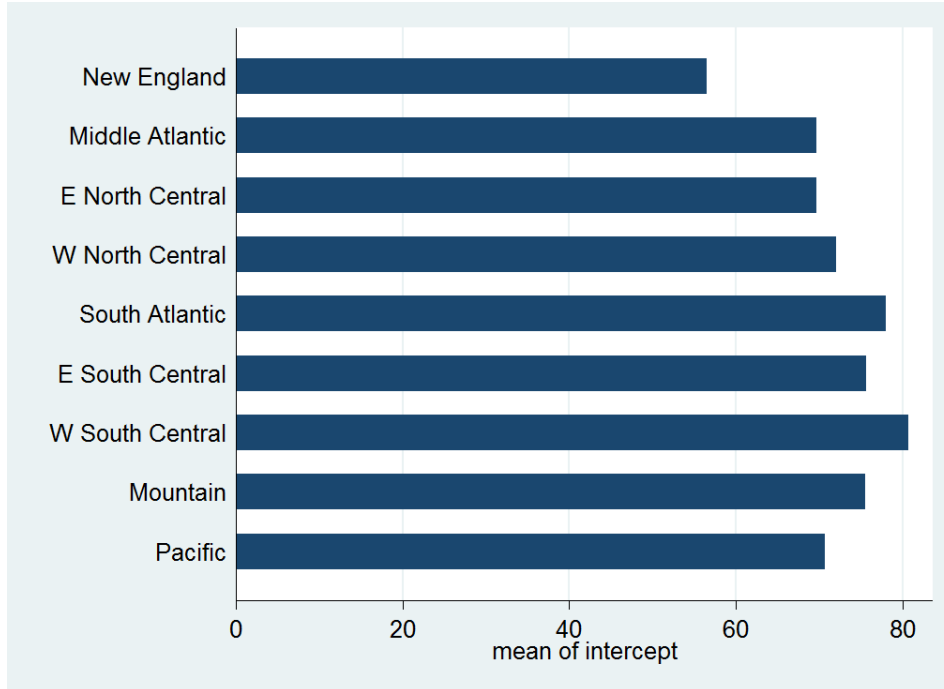
Ave of $Y_i = 42.38305$

$u = 58.23352$

BLUP of $b_i = -15.22133$

Random Intercept Model

```
. xtmixed bush logdens minority colled || cendiv:  
. predict randint0, reffects  
. gen intercept=_b[_cons]+randint0  
. graph hbar (mean) intercept, over(cendiv)
```



$$(2.2) \quad y_{ij} = \beta_0 + u_{0j} + \beta_1 X_{1,ij} + \beta_2 X_{2,ij} + \beta_3 X_{3,ij} + \varepsilon_{ij}$$
$$= \beta_{0j} + \beta_1 X_{1,ij} + \beta_2 X_{2,ij} + \beta_3 X_{3,ij} + \varepsilon_{ij} .$$

- Total effects of intercepts=fixed effects + predicted random effects
- $E(y_{ij}|x_{1,ij}, x_{2,ij}, x_{3,ij}) = \beta_0 + E(u_{0j}) + \text{constant}$, for a given j .
- So the "randint0" gives the difference of the intercepts from the fixed effect intercept (population intercept) for each census division
- Interpretation: at any given level of the 3 predictors (logdens, minority and colled), percentage of voters for Bush is on average 16 points lower in the New England (NE) region than in the West/North/Central region, and 22 percent lower in the NE region than in the South Atlantic region.

Random Regression Model: Intercept+Slope

```
. xtmixed bush logdens minority colled || cendiv: logdens
```

Performing EM optimization:

Performing gradient-based optimization:

Iteration 0: log likelihood = -11298.734

Iteration 1: log likelihood = -11298.734

Computing standard errors:

Mixed-effects ML regression

Group variable: cendiv

Number of obs = 3041

Number of groups = 9

Obs per group: min = 67

avg = 337.9

max = 616

Wald chi2(3) = 806.25

Prob > chi2 = 0.0000

Log likelihood = -11298.734

bush	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
logdens	-3.310313	1.114965	-2.97	0.003	-5.495605	-1.125021
minority	-.3616886	.0130709	-27.67	0.000	-.387307	-.3360702
colled	-.1173469	.0360906	-3.25	0.001	-.1880833	-.0466105
_cons	70.12095	2.955209	23.73	0.000	64.32885	75.91305

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]	
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cendiv: Independent

sd(logdens)	3.113575	.8143897	1.86474	5.198768
-------------	----------	----------	---------	----------

sd(_cons)	8.5913	2.232214	5.162945	14.29619
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sd(Residual)	9.825565	.1264176	9.580889	10.07649
--------------	----------	----------	----------	----------

LR test vs. linear regression: chi2(2) = 538.10 Prob > chi2 = 0.0000

Note: LR test is conservative and provided only for reference.

□ Compare this model with the RIM

$$\begin{aligned}
 (2.5) \quad y_{ij} &= \beta_0 + u_{0j} + \beta_1 X_{1,ij} + u_{1j} X_{1,ij} + \beta_2 X_{2,ij} + \beta_3 X_{3,ij} + \varepsilon_{ij} \\
 &= \beta_{0j} + \beta_{1j} X_{1,ij} + \beta_2 X_{2,ij} + \beta_3 X_{3,ij} + \varepsilon_{ij} .
 \end{aligned}$$

Random Regression Model: Intercept+Slope

▣ Compare this model with the RIM

```
. quietly xtmixed bush logdens minority colled || cendiv:  
. estimates store m1  
  
. quietly xtmixed bush logdens minority colled || cendiv: logdens  
. estimates store m2  
  
. lrtest m1 m2
```

```
Likelihood-ratio test                                LR chi2(1)  =      82.11  
(Assumption: m1 nested in m2)                       Prob > chi2 =     0.0000
```

Note: The reported degrees of freedom assumes the null hypothesis is not on the boundary of the parameter space. If this is not true, then the reported test is conservative.

▣ Inclusion of random slope into the model brought significant improvement in it.

Random Regression Model: Intercept+Slope

- Are the random intercept and slope are correlated?

```
. xtmixed bush logdens minority collid || cendiv: logdens,  
cov(unstructured)
```

- The random intercept and slope are correlated

- So we include the correlation.

Performing EM optimization:

Performing gradient-based optimization:

Iteration 0: log likelihood = -11296.31
Iteration 1: log likelihood = -11296.31

Computing standard errors:
Mixed-effects ML regression
Group variable: cendiv

Number of obs = 3041
Number of groups = 9
Obs per group: min = 67
avg = 337.9
max = 616
Wald chi2(3) = 799.68
Prob > chi2 = 0.0000

Log likelihood = -11296.31

bush	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
logdens	-3.150009	1.169325	-2.69	0.007	-5.441844 - .8581751
minority	-.3611161	.0130977	-27.57	0.000	-.3867872 - .3354451
collid	-.1230445	.0361363	-3.41	0.001	-.1938704 -.0522186
_cons	69.85194	3.168479	22.05	0.000	63.64184 76.06204

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]
cendiv: Unstructured			
sd(logdens)	3.282749	.8547255	1.970658 5.468447
sd(_cons)	9.240389	2.402183	5.551459 15.3806
corr(logdens, _cons)	-.675152	.1958924	-.9096869 -.1140963

	sd(Residual)	Estimate	Std. Err.	[95% Conf. Interval]
		9.823658	.1263468	9.579118 10.07444

LR test vs. linear regression: chi2(3) = 542.95 Prob > chi2 = 0.0000

Note: LR test is conservative and provided only for reference.

```
. lrtest m1 m2
```

Likelihood-ratio test
(Assumption: m1 nested in m2)

LR chi2(1) = 4.44
Prob > chi2 = 0.0352

Random Regression Model: Intercept+Slope

- Predict intercepts and slopes within each of the 9 regions

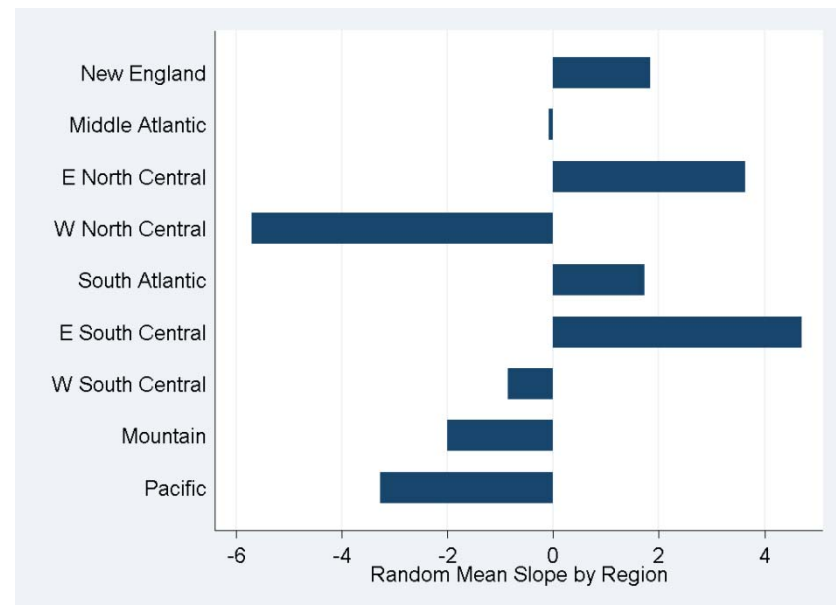
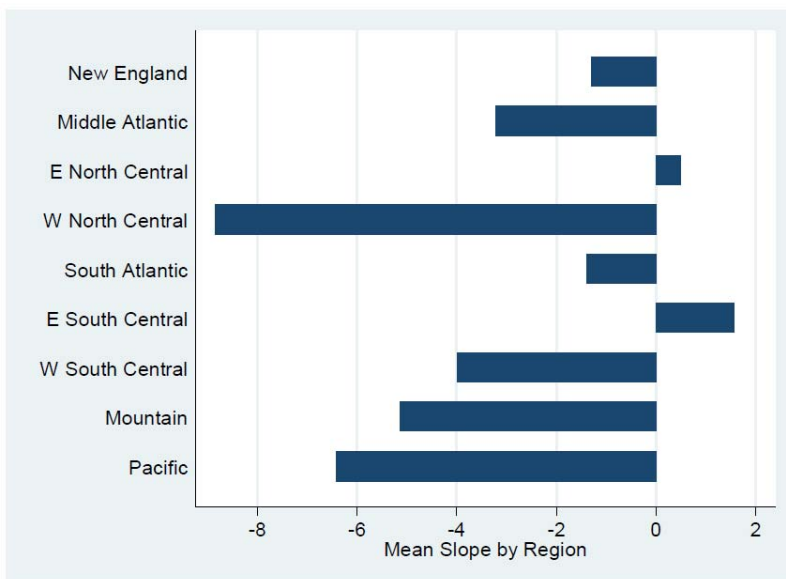
```
. predict randslo1 randint1, reffects  
. table cendiv, contents (mean randslo1 mean randint1)
```

```
-----  
Census division |  
(9)             | mean(randslo1)  mean(randint1)  
-----+-----  
    New England |          1.760172      -18.79535  
Middle Atlantic |         -.0611391       -2.198096  
E North Central |          3.618166       -9.007434  
W North Central |         -5.65562        8.328082  
  South Atlantic |          1.738479        2.875959  
E South Central |          4.633666       -4.289312  
W South Central |         -.8330051        10.8619  
      Mountain |         -1.975749        7.232863  
        Pacific |         -3.224968        4.991385  
-----
```

Random Regression Model: Intercept+Slope

- For each region, the mean slope is the fixed effect slope + the mean in that region of the pertinent random slopes: combined slope

```
. gen slope1 = randslo1 + _b[logdens]  
. graph hbar (mean) slope1, over(cendiv) ytitle ("Mean Slope by  
Region")
```



Random Regression Model: Multiple Random Slopes

```
□ . xtmixed bush logdens minority colled || cendiv: logdens  
   minority colled
```

Performing EM optimization:

Performing gradient-based optimization:

Iteration 0: log likelihood = -11184.804
Iteration 1: log likelihood = -11184.804

Computing standard errors:

Mixed-effects ML regression	Number of obs	=	3041
Group variable: cendiv	Number of groups	=	9
	Obs per group: min	=	67
	avg	=	337.9
	max	=	616
	Wald chi2(3)	=	52.49
	Prob > chi2	=	0.0000

Log likelihood = -11184.804

bush	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
logdens	-2.717128	1.373684	-1.98	0.048	-5.409499 -.0247572
minority	-.3795605	.0560052	-6.78	0.000	-.4893286 -.2697924
colled	-.1707863	.1727742	-0.99	0.323	-.5094175 .167845
_cons	70.86653	3.435918	20.63	0.000	64.13225 77.6008

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]
cendiv: Independent			
sd(logdens)	3.868421	.9832899	2.350564 6.366421
sd(minority)	.153172	.0439569	.0872777 .2688161
sd(colled)	.5032414	.1241234	.310334 .8160625
sd(_cons)	10.01157	2.547813	6.079707 16.48625

sd(Residual)	9.375994	.1209753	9.141859 9.616124
--------------	----------	----------	-------------------

LR test vs. linear regression: chi2(4) = 765.96 Prob > chi2 = 0.0000

Note: LR test is conservative and provided only for reference.

Random Regression Model: Multiple Random Slopes

```
. estimates store full  
. quietly xtmixed bush logdens minority colled || cendiv: logdens minority  
. estimates store nocolled  
. lrtest nocolled full
```

Likelihood-ratio test	LR chi2(1) =	197.33
(Assumption: nocolled nested in full)	Prob > chi2 =	0.0000

- ▣ This shows that the full model fits significantly better
- ▣ And thus there's no evidence in the data set that would warrant excluding the random effect associated with the variable *colled*.

Fixed, Random and Total Effects

- Random effects help better fit models since they represent *heterogeneity* in the data
- Oftentimes, random effects are considered not of substantive interest
- In some empirical settings, however, one may be interested in the random effects themselves.
- The latter usually happens when of concern are the total effects, which are defined as the sum of random and fixed effects for a given predictor.

```
. quietly xtmixed bush logdens minority colled || cendiv: logdens minority colled  
. predict relogdens reminority recolled re_cons, reffects  
. d relogdens-re_cons
```

variable name	storage type	display format	value label	variable label
relogdens	float	%9.0g		BLUP r.e. for cendiv: logdens
reminority	float	%9.0g		BLUP r.e. for cendiv: minority
recolled	float	%9.0g		BLUP r.e. for cendiv: colled
re_cons	float	%9.0g		BLUP r.e. for cendiv: _cons

Fixed, Random and Total Effects

- Computing total effects, for example *colled*:

```
. gen tecolled = recolled + _b[colled]

. label variable tecolled "Random Plus Fixed Effect of Coll_Ed"

. d tecolled
```

variable name	storage type	display format	value label	variable label
tecolled	float	%9.0g		Random Plus Fixed Effect of Coll_Ed

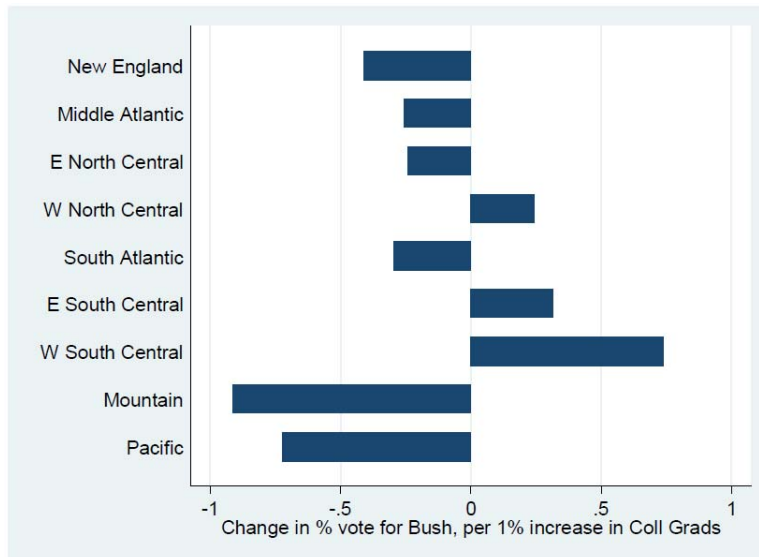
- Means for the random effects and total effects of *colled* per geographic region

```
. table cendiv, contents(mean recolled mean tecolled)
```

Census division (9)	mean(recolled)	mean(tecolled)
New England	-.2406214	-.4122202
Middle Atlantic	-.0856313	-.2572301
E North Central	-.0695729	-.2411717
W North Central	.4153244	.2437256
South Atlantic	-.125596	-.2971948
E South Central	.4871398	.315541
W South Central	.9101545	.7385557
Mountain	-.7414296	-.9130284
Pacific	-.5497676	-.7213664

Fixed, Random and Total Effects

```
. graph hbar (mean) tecolled, over(cendiv) ytitle ("Change in %  
vote for Bush, per 1% increase in Coll Grads")
```



- A model with a random effect for *colled* was better than the one without.
- In the fixed effect model, the various random effects across region will be averaged out into a single fixed effect.
- The total effects of *colled* range from substantially negative to substantially positive, will be lost.
- In fact, the fixed effect estimate for *colled* is -0.17 if using conventional regression, which would be quite misleading.

Nested Levels

- Is it meaning for to consider counties nested perhaps also in states?
- Consider variable *colled* nesting within states.

```
. xtmixed bush logdens minority colled || cendiv: logdens  
minority colled || state: colled
```

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]	
-----+-----				
cendiv: Independent				
sd(logdens)	2.703845	.7727275	1.544252	4.734186
sd(minority)	.1465435	.0428326	.0826365	.2598728
sd(colled)	.3683903	.0962733	.220729	.6148326
sd(_cons)	8.416873	2.417524	4.793643	14.77869
-----+-----				
state: Independent				
sd(colled)	.1305727	.039009	.0727032	.2345047
sd(_cons)	5.883451	.7431715	4.593166	7.536196
-----+-----				
sd(Residual)	7.863302	.1027691	7.664436	8.067328
-----+-----				
LR test vs. linear regression:	chi2(6) = 1695.92		Prob > chi2 = 0.0000	

Note: LR test is conservative and provided only for reference.

- All the random effects are significant

Nested Levels

```
. estimates store state  
. lrtest full state
```

Likelihood-ratio test

(Assumption: full nested in state)

LR chi2(2) = 929.97

Prob > chi2 = 0.0000

- The “state” model fits significantly better

```
. predict re*, reffects
```

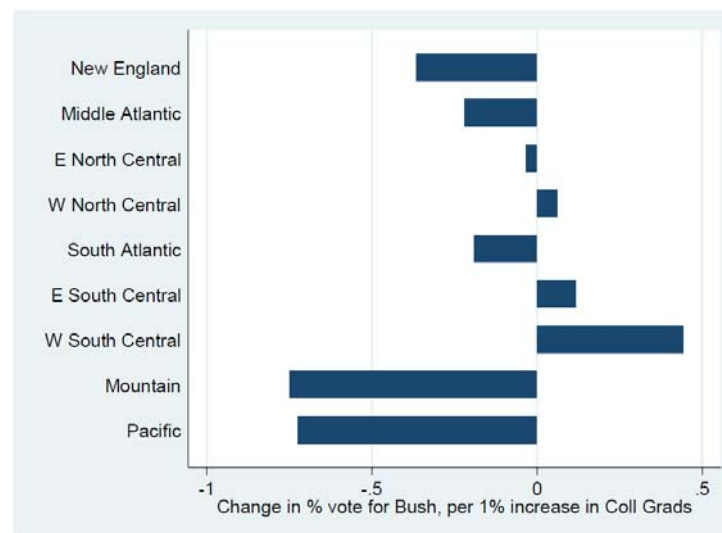
```
re1          float    %9.0g  
re2          float    %9.0g  
re3          float    %9.0g  
re4          float    %9.0g  
re5          float    %9.0g  
re6          float    %9.0g
```

```
BLUP r.e. for cendiv: logdens  
BLUP r.e. for cendiv: minority  
BLUP r.e. for cendiv: colled  
BLUP r.e. for cendiv: _cons  
BLUP r.e. for state: colled  
BLUP r.e. for state: _cons
```

- Total effects for *colled*: $\text{tecolled2} = \text{re3} + \text{re5} + _b[\text{colled}]$

```
. graph hbar (mean) tecolled2, over(cendiv) ytitle ("Change in % vote for  
Bush, per 1% increase in Coll Grads")
```

- The central regions have now somewhat closer to 0 (in absolute value)
- That for mountain regions marginally less negative
- Accounting for the intermediate level of nesting is associated with less pronounced effects of *colled* upon the response variable in those regions



Binary Responses

- In social and behavioral research, there is the case that examining relationships between discrete response and explanatory variables.

```
. use c:\stata\data\gss.dta, clear
(General Social Survey 2006)
```

```
Contains data from c:\stata\data\GSS_SwS1.dta
  obs:      1,595                      General Social Survey 2006
  vars:      17                      1 May 2011 13:45
  size:     54,230 (99.5% of memory free)
```

storage	display	value		
variable name	type	format	label	variable label
id	int	%9.0g		Case number (Statistics
w/Stata)				
finalwt	float	%9.0g		weight variable -- wtssall
cendiv	byte	%12.0g	cendiv	Census division (9)
age	byte	%8.0g	age	Age in years
educ	byte	%8.0g	educ	Highest year of schooling
completed				
gender	byte	%9.0g	gender	Respondent gender
income	float	%9.0g	income	Family income in constant
dollars				
loginc	float	%9.0g		log10(family income, +1)
logsize	float	%9.0g		log10(size of place in
thousands, +1)				
married	byte	%9.0g	married	Married or unmarried?
minority	byte	%11.0g	minority	Minority status
politics	byte	%20.0g	politics	Think of self as liberal or
conservative?				
bush	byte	%11.0g	bush	Voted for GW Bush in 2004?
grass	byte	%9.0g	grass	Should marijuana be made legal?
gunlaw	byte	%8.0g	gunlaw	Favor or oppose gun permits?
postlife	byte	%8.0g	postlife	Believe in life after death?
enviro	byte	%11.0g	enviro	Govt spending on environment?

```
Sorted by: id
```

Binary Responses

- Overall distributed votes

```
. tab bush
Voted for |
  GW Bush in |
      2004? |      Freq.      Percent      Cum.
-----+-----
Kerry/Nader |      776      48.65      48.65
      Bush |      819      51.35     100.00
-----+-----
      Total |     1,595     100.00
```

- We use *xtnlogit*
- Consider the log-odds of voting for Bush here.

Random Intercept Model

- Random intercept model

$$(3.1) \quad \ln[p_{ij}/(1-p_{ij})] = \beta_0 + u_{0j} + \beta_1 X_{1,ij} + \beta_2 X_{2,ij} + \beta_3 X_{3,ij}.$$

- There is no error term: this is a model stated already in terms of the expectation of 'event'.

`. xtlogit bush logsize minority educ || cendiv:`

- The last line suggests that inclusion of a random intercept improves the model over the fixed effects only model (conventional logistic regression)

- We don't have a residual random effect

Refining starting values:

```
Iteration 0: log likelihood = -1008.5032 (not concave)
Iteration 1: log likelihood = -1001.8839
Iteration 2: log likelihood = -1001.4243
```

Performing gradient-based optimization:

```
Iteration 0: log likelihood = -1001.4243
Iteration 1: log likelihood = -1001.4226
Iteration 2: log likelihood = -1001.4226
```

Mixed-effects logistic regression
Group variable: cendiv

```
Number of obs      =    1595
Number of groups   =      9
Obs per group: min =     62
                  avg =   177.2
                  max =    336
Wald chi2(3)       =   156.51
Prob > chi2        =    0.0000
```

Integration points = 7
Log likelihood = -1001.4226

	bush	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
logsize		-.2456232	.0723751	-3.39	0.001	-.3874758	-.1037707
minority		-1.988091	.1693794	-11.74	0.000	-2.320069	-1.656114
educ		-.0435621	.020059	-2.17	0.030	-.082877	-.0042472
_cons		1.331863	.3044779	4.37	0.000	.7350978	1.928629

Random-effects Parameters		Estimate	Std. Err.	[95% Conf. Interval]	
cendiv: Identity					
	sd(_cons)	.2631985	.0937197	.1309747	.5289069

LR test vs. logistic regression: chibar2(01) = 9.78 Prob>=chibar2 = 0.0009

Random Regression Model

- Will it be better to include a random slope for *logsize*?

$$(3.2) \quad \ln[p_{ij}/(1-p_{ij})] = (\beta_0 + u_{0j}) + (\beta_1 + u_{1j})X_{1,ij} + \beta_2 X_{2,ij} + \beta_3 X_{3,ij}.$$

```
. xtmelogit bush logsize minority educ
    || cendiv: logsize, covariance(unstructured)
. estimates store a
```

- The random intercepts are perhaps not significant as an effect
- They are not really related to random slopes

Refining starting values:

```
Iteration 0: log likelihood = -1006.9894 (not concave)
Iteration 1: log likelihood = -995.258
Iteration 2: log likelihood = -993.39684
```

Performing gradient-based optimization:

```
Iteration 0: log likelihood = -993.39684 (not concave)
Iteration 1: log likelihood = -993.29552
Iteration 2: log likelihood = -993.19391
Iteration 3: log likelihood = -993.18826
Iteration 4: log likelihood = -993.18808
Iteration 5: log likelihood = -993.18808
```

Mixed-effects logistic regression	Number of obs	=	1595
Group variable: cendiv	Number of groups	=	9
	Obs per group: min	=	62
	avg	=	177.2
	max	=	336
Integration points = 7	Wald chi2(3)	=	136.94
Log likelihood = -993.18808	Prob > chi2	=	0.0000

bush	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
logsize	-.3067214	.1351804	-2.27	0.023	-.57167 -.0417728
minority	-1.929399	.1705009	-11.32	0.000	-2.263574 -1.595223
educ	-.041411	.0201692	-2.05	0.040	-.0809419 -.00188
_cons	1.361903	.3014102	4.52	0.000	.7711501 1.952656

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]
cendiv: Unstructured			
sd(logsize)	.3249856	.1153993	.1620356 .6518051
sd(_cons)	.2166197	.1880638	.0395099 1.187653
corr(logsize,_cons)	-.8708498	.2677054	-.9982073 .6831364

LR test vs. logistic regression: chi2(3) = 26.25 Prob > chi2 = 0.0000

Random Slope Only Model

- Dispense with the random intercepts

$$(3.3) \quad \ln[p_{ij}/(1-p_{ij})] = \beta_0 + (\beta_1 + u_{1j})X_{1,ij} + \beta_2 X_{2,ij} + \beta_3 X_{3,ij}.$$

```
. xtmelogit bush logsize minority educ
  || cendiv: logsize, nocons
. estimates store b
. lrtest b a
```

- The model (3.3), the random-slope-only (RSO) is not significantly worse than the random regression (RR) Model (3.2)

- The RSO model is preferable

Mixed-effects logistic regression
Group variable: cendiv

Number of obs = 1595
Number of groups = 9
Obs per group: min = 62
 avg = 177.2
 max = 336

Integration points = 7
Log likelihood = -994.02453

Wald chi2(3) = 150.81
Prob > chi2 = 0.0000

bush	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
logsize	-.2988801	.1097208	-2.72	0.006	-.5139289	-.0838312
minority	-1.969727	.1666701	-11.82	0.000	-2.296395	-1.64306
educ	-.0414891	.0201459	-2.06	0.039	-.0809745	-.0020038
_cons	1.361989	.2909969	4.68	0.000	.7916459	1.932333

Random-effects Parameters		Estimate	Std. Err.	[95% Conf. Interval]	
cendiv: Identity					
sd(logsize)		.2340086	.071626	.128438	.4263536

LR test vs. logistic regression: chibar2(01) = 24.58 Prob>=chibar2 = 0.0000

Likelihood-ratio test
(Assumption: b nested in a)

LR chi2(2) = 1.67
Prob > chi2 = 0.4332

Model Choice

- Compare RSO (random-slope-only) with RIM (random intercept model): they are not nested, can not use likelihood ratio test
- Akaike's Information Criterion (AIC)
(3.4) **$AIC = -2 * \log \text{likelihood} + 2 * \# \text{parameters}$**
- Log-likelihood values are from the Stata pertinent model outputs
- Number of parameters can be counted:
 - RIM model (3.1) has 5 parameters (4 fixed and 1 random)
 - RSO model (3.3) has 5 parameters (4 fixed and 1 random)
- `gen AIC_RIM=-2*(-1001.4226)+2*5`
- `gen AIC_RSO=-2*(-994.02453)+2*5`
- `AIC_RIM=2012.845`
- `AIC_RSO=1998.049`

- RSO has a lower AIC, so it is preferable to the RIM
- The decision suggests that the regional variability in the pattern of voting for Bush is best modeled as variations in the effect of logsize (an indicator of urbanization or urban-ness) upon the odds of such vote, after accounting for minority and educational levels

Day 2 Conclusion

- ▣ This part was devoted to intermediate and some advanced topics within multilevel modeling
- ▣ Issues pertaining to the choice between single level statistical models and two level models, as well as between two-level and three-level models, are of special relevance in empirical behavioral, social and biomedical research.
- ▣ These issues were attended to from a model fitting perspective, in the context of mixed models
- ▣ Mixed modeling issues pertaining to random effect and total effect estimation were also discussed

The End

