Introduction to Bayesian Data Analysis and Markov Chain Monte Carlo

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Abstract

The purpose of this talk is to give a brief overview of Bayesian Inference and Markov Chain Monte Carlo methods, including the Gibbs Sampler and Metropolis Hastings algorithm.

Outline

- Bayesian vs. Frequentist paradigm
- Bayesian Inference and MCMC
 - ⋆ Gibbs Sampler
 - ⋆ Metropolis-Hastings Algorithm
- Assessing Convergence of MCMC
- Hierarchical Model Example
- MCMC: Benefits and Cautions

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MCMC OVERVIEW

Frequentist vs. Bayesian paradigms

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To a Bayesian:

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- \star Inference is performed via the posterior distribution $f(\mathbf{\Theta}|\mathbf{X})$.
- ★ We <u>can</u> make probability statements about parameters, since they are random quantities (e.g. credible intervals)

$$f(\mathbf{\Theta}|\mathbf{X}) = \frac{f(\mathbf{X}|\mathbf{\Theta})f(\mathbf{\Theta})}{f(\mathbf{X})}$$

• The posterior distribution is computed by applying **Bayes' Rule:**

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- End result: A (correlated) sample from the stationary distribution.

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- Samples can be used to perform any Bayesian inference of interest.
- How do we generate the Markov Chain?

9

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Conditions on:

- ★ The data X
- \star The values for all other parameters Θ_{-i} .

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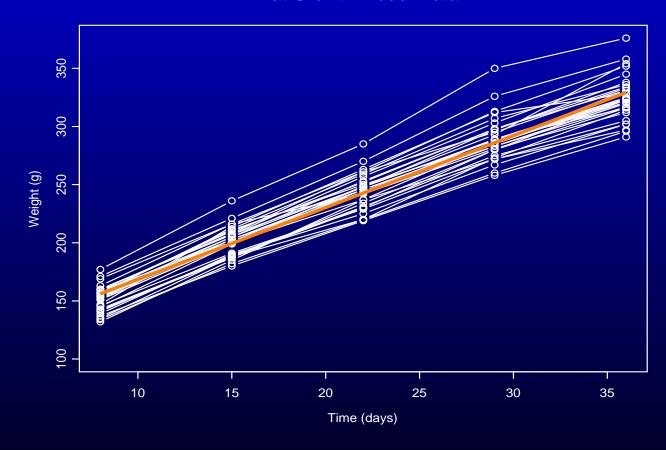
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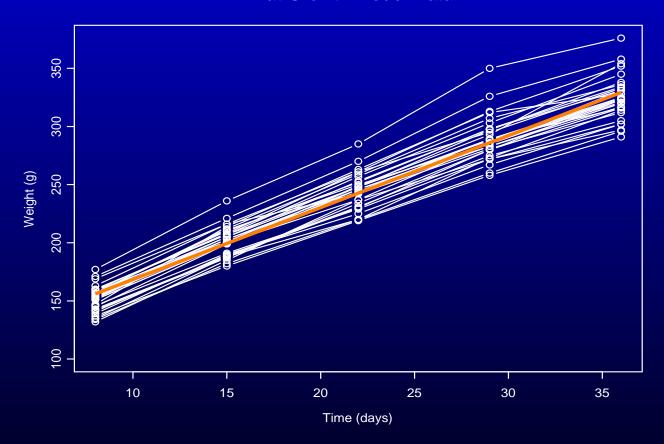
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 Can estimate mean growth curve by linear regression, but growth curve models necessary to get standard errors right.

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Conclusions

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- Other book: Gelman, Carlin, Stern, & Rubin (1995) Bayesian Data Analysis

MCMC OVERVIEW 16

References

Gelman A and Rubin DB (1992). Inference from iterative simulation using multiple sequences. *Statistical Science* **7**, 457 75472.

Geman S and Geman D (1984). Stochastic relaxation, Gibbs distributions, and the Bayesian restoration of images. *IEEE Trans. Pattn. Anal. Mach. Intel.* **6**, 721 75741.

Geweke J (1992). Evaluation of accuracy of sampling-based approaches to the calculation of posterior moments. In *Bayesian Statistics 4*(ed. JM Bernardo, J Berger, AP Dawid and AFM Smith), pp. 169 75193. Oxford University Press.

Gilks WR, Richardson S, and Spiegelhalter DJ (1996). Markov Chain Monte Carlo in Practice, Chapman and Hall.

Hastings WK (1970). Monte Carlo sampling methods using Markov chains and their applications. *Biometrika* **57**, 97 75109.

Metropolis N, Rosenbluth AW, Rosenbluth MN, Teller AH and Teller E (1953). Equations of state calculations by fast computing machine. *J. Chem. Phys.* 21, 1087 751091.